

Notions on optimal transport

Following (1), let us take $\mathcal{P}(\Omega)$, the space of probability distributions on Ω . For μ, ν in $\mathcal{P}(\Omega)$, let us define $\Pi(\mu, \nu)$ the set of all probability measures π on $\Omega \times \Omega$ with first marginal μ and second marginal ν . The optimal transport cost between the two measures is defined as

$$C(\mu, \nu) = \inf_{\pi \in \Pi(\mu, \nu)} \int c(x, y) d\pi(x, y) \quad (1)$$

where $c(x, y)$ is the cost of transporting one unit of mass from x to y . A probability π that achieves the minimum in (1) is called an optimal coupling, with an associated random variable (X, Y) that has joint distribution π . When μ and ν are discrete, i.e., $\mu = \sum_{i=1}^n p_i \delta_{x_i}$ and $\nu = \sum_{j=1}^m q_j \delta_{y_j}$, with $x_i, y_j \in \mathbb{R}^d$, the optimal transport problem can be solved as a linear program (see (2)) where

$$C(\mu, \nu) = \sum_{i=1}^n \sum_{j=1}^m w_{ij}^* c(x_i, y_j),$$

and (w_{ij}^*) are the solutions of the optimal transport linear program

$$\begin{aligned} & \text{minimize} && \sum_{i=1}^n \sum_{j=1}^m w_{ij} c(x_i, y_j) \\ & \text{subject to} && w_{ij} \geq 0, && 1 \leq i \leq n, 1 \leq j \leq m \\ & && \sum_{j=1}^m w_{ij} = p_i, && 1 \leq i \leq n \\ & && \sum_{i=1}^n w_{ij} = q_j, && 1 \leq j \leq m \\ & && \sum_{i=1}^n \sum_{j=1}^m w_{ij} = 1. \end{aligned}$$

For $(\Omega = \mathbb{R}^d, \|\cdot\|)$, with $\|\cdot\|$ the Euclidean norm, and $p \in [1, \infty)$, the p -Wasserstein distance between μ and ν is defined as

$$\begin{aligned} \mathcal{W}_p^p(\mu, \nu) &= \inf_{\pi \in \Pi(\mu, \nu)} \int \|x - y\|^p d\pi(x, y) \\ &= \inf \{E\|X - Y\|^p, \mathcal{L}(X) = \mu, \mathcal{L}(Y) = \nu\}, \end{aligned}$$

where $\mathcal{L}(X)$ refers to the law of X .

We present the entropy regularized Wasserstein distance, since it is strictly convex and there are efficient solutions based on the Sinkhorn algorithm (see (3)). For a fixed $\gamma > 0$ the regularized Wasserstein distance is defined as

$$\mathcal{W}_\gamma(\mu, \nu) = \sum_{i=1}^n \sum_{j=1}^m w_{ij}^* \|x_i - y_j\|^2 + \gamma \sum_{i=1}^n \sum_{j=1}^m w_{ij}^* \log w_{ij}^*, \quad (2)$$

where (w_{ij}^*) are the solutions of the optimal transport linear program

$$\begin{aligned} & \text{minimize} && \sum_{i=1}^n \sum_{j=1}^m w_{ij} \|x_i - y_j\|^2 + \gamma \sum_{i=1}^n \sum_{j=1}^m w_{ij} \log w_{ij} \\ & \text{subject to} && w_{ij} \geq 0, && 1 \leq i \leq n, 1 \leq j \leq m \\ & && \sum_{j=1}^m w_{ij} = p_i, && 1 \leq i \leq n \\ & && \sum_{i=1}^n w_{ij} = q_j, && 1 \leq j \leq m \\ & && \sum_{i=1}^n \sum_{j=1}^m w_{ij} = 1. \end{aligned}$$

Let us denote $\mathcal{P}_2(\mathbb{R}^d)$ the set of probability measures on \mathbb{R}^d with finite second moment and let us consider $\mathcal{W}_2(\mu, \nu)$ for $\mu, \nu \in \mathcal{P}(\mathbb{R}^d)$. In (4) the notions of k -barycenter and trimmed k -barycenter were introduced, building on the concept of Wasserstein barycenter introduced in

(5; 6). A k -barycenter of probabilities $\{\mu_1, \dots, \mu_n\}$ in $\mathcal{P}_2(\mathbb{R}^d)$ with weights $\lambda_1, \dots, \lambda_n$ is any k -set $\{\bar{\mu}_1, \dots, \bar{\mu}_k\}$ in $\mathcal{P}_2(\mathbb{R}^d)$ such that for any $\{\nu_1, \dots, \nu_k\} \subset \mathcal{P}_2(\mathbb{R}^d)$ we have that

$$\sum_{i=1}^n \lambda_i \min_{j \in \{1, \dots, k\}} \mathcal{W}_2^2(\mu_i, \bar{\mu}_j) \leq \sum_{i=1}^n \lambda_i \min_{j \in \{1, \dots, k\}} \mathcal{W}_2^2(\mu_i, \nu_j). \quad (3)$$

An α -trimmed k -barycenter of $\{\mu_1, \dots, \mu_n\}$ with weights as before is any k -set $\{\bar{\mu}_1, \dots, \bar{\mu}_k\}$ with weights $\bar{\lambda} = (\bar{\lambda}_1, \dots, \bar{\lambda}_n) \in \Lambda_\alpha(\lambda)$ such that

$$\sum_{i=1}^n \bar{\lambda}_i \min_{j \in \{1, \dots, k\}} \mathcal{W}_2^2(\mu_i, \bar{\mu}_j) = \min_{\{\nu_1, \dots, \nu_k\} \subset \mathcal{P}_2(\mathbb{R}^d), \lambda^* \in \Lambda_\alpha(\lambda)} \sum_{i=1}^n \lambda_i^* \min_{j \in \{1, \dots, k\}} \mathcal{W}_2^2(\mu_i, \nu_j), \quad (4)$$

where $\Lambda_\alpha(\lambda) = \{\lambda^* = (\lambda_1^*, \dots, \lambda_n^*) : 0 \leq \lambda_i^* \leq \lambda_i / (1 - \alpha), \sum_{i=1}^n \lambda_i^* = 1\}$.

Broadly speaking k -barycenters can be thought of as an extension of k -means to the space of probabilities with finite second order, since we can rewrite (3) as

$$\min_{\mathfrak{S}} \sum_{j=1}^k \sum_{\mu_i \in \mathfrak{S}_j} \lambda_i \mathcal{W}_2^2(\mu_i, \bar{\mu}_j) \quad (5)$$

where $\mathfrak{S} = \{\mathfrak{S}_1, \dots, \mathfrak{S}_k\}$ is a partition of $\{\mu_1, \dots, \mu_n\}$ and $\bar{\mu}_j$ is the barycenter of the elements in \mathfrak{S}_j . Therefore, trimmed k -barycenters may be matched to trimmed k -means. As stated in (4), efficient computations can be done when dealing with location-scatter families of absolutely continuous distributions in $\mathcal{P}_2(\mathbb{R}^d)$. A notable example being the family of multivariate Gaussian distributions.

References

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